

Union-Intersection Based Goodness-of-Fit Tests in Terms of Local Levels

V. Gontscharuk and H. Finner

Institute for Biometrics and Epidemiology,
German Diabetes Center at the Heinrich-Heine-University Düsseldorf,
Leibniz Center for Diabetes Research

Multiple Tests and Confidence Intervals
Cologne, 24. June 2015

Goodness-of-Fit Tests

$X_1, \dots, X_n \stackrel{iid}{\sim} F$ continuous, $X_{1:n}, \dots, X_{n:n}$ related order statistics

$$H_0^+ : F(x) \leq F_0(x) \text{ or } H_0 : F(x) = F_0(x)$$

φ^+ rejects H_0^+ (i.e., $\varphi^+ = 1$) iff $X_{i:n} \leq c_{i,n}$ for some i .

φ rejects H_0 (i.e., $\varphi = 1$) iff $X_{i:n} \leq c_{i,n}$ or $X_{i:n} \geq \tilde{c}_{i,n}$ for some i .

Assumption: $0 \leq c_{1,n} < \dots < c_{n,n} < 1$, $0 < \tilde{c}_{1,n} < \dots < \tilde{c}_{n,n} \leq 1$
and $c_{i,n} < \tilde{c}_{i,n}$, $i = 1, \dots, n$.

Examples: Kolmogorov-Smirnov, Anderson-Darling,
Berk-Jones tests and other tests based on φ -divergences

Goodness-of-Fit Tests

$X_1, \dots, X_n \stackrel{iid}{\sim} F$ continuous, $X_{1:n}, \dots, X_{n:n}$ related order statistics

$$H_0^+ : F(x) \leq x \text{ or } H_0 : F(x) = x$$

φ^+ rejects H_0^+ (i.e., $\varphi^+ = 1$) iff $X_{i:n} \leq c_{i,n}$ for some i .

φ rejects H_0 (i.e., $\varphi = 1$) iff $X_{i:n} \leq c_{i,n}$ or $X_{i:n} \geq \tilde{c}_{i,n}$ for some i .

Assumption: $0 \leq c_{1,n} < \dots < c_{n,n} < 1$, $0 < \tilde{c}_{1,n} < \dots < \tilde{c}_{n,n} \leq 1$
and $c_{i,n} < \tilde{c}_{i,n}$, $i = 1, \dots, n$.

Examples: Kolmogorov-Smirnov, Anderson-Darling,
Berk-Jones tests and other tests based on φ -divergences

The Union-Intersection Principle and Local Levels

$U_1, \dots, U_n \stackrel{iid}{\sim} U(0, 1)$, $U_{1:n}, \dots, U_{n:n}$ related order statistics

H_i^+ (or H_i) is true if $X_{i:n} \stackrel{st.}{\geq} U_{i:n}$ (or $X_{i:n} \stackrel{D}{=} U_{i:n}$)

$\Rightarrow H_0^+ \subseteq \bigcap_{i=1}^n H_i^+$ and $H_0 \subseteq \bigcap_{i=1}^n H_i$

φ_i^+ rejects H_i^+ (i.e., $\varphi_i^+ = 1$) iff $X_{i:n} \leq c_{i,n}$.

φ_i rejects H_i (i.e., $\varphi_i = 1$) iff $X_{i:n} \leq c_{i,n}$ or $X_{i:n} \geq \tilde{c}_{i,n}$.

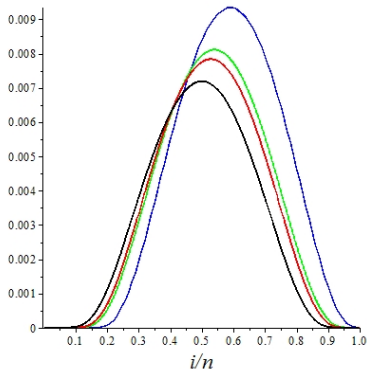
$\Rightarrow \varphi^+ = \max_i \varphi_i^+$ and $\varphi = \max_i \varphi_i$ are union-intersection tests.

Local levels:

$$\alpha_{i,n}^+ = \mathbb{P}(\varphi_i^+ = 1 | H_0) = \mathbb{P}(U_{i:n} \leq c_{i,n})$$

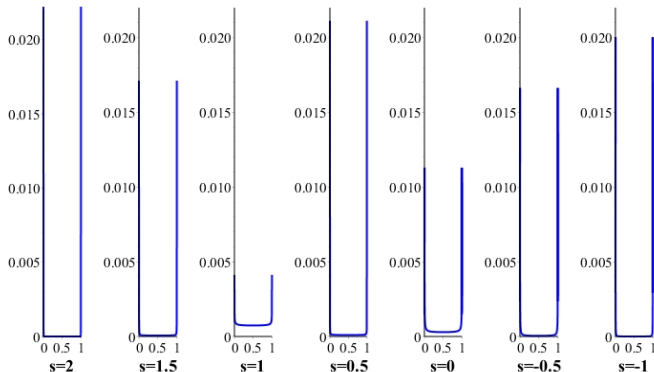
$$\alpha_{i,n} = \mathbb{P}(\varphi_i = 1 | H_0) = \mathbb{P}(U_{i:n} \leq c_{i,n}) + \mathbb{P}(U_{i:n} \geq \tilde{c}_{i,n})$$

Local Levels of Kolmogorov-Smirnov Tests



$\alpha_{i,n}^+$ for $\alpha = 0.05$ and $n = 100, 500, 1000$ and the corresponding asymptotic local levels (black curve)

Local Levels of Tests Based on φ -Divergences



$\alpha_{i,n} \equiv \alpha_{i,n}(s)$ with $\alpha = 0.05$ and $n = 1000$

$s = 2$: Higher Criticism (HC), $s = 1$: Berk-Jones (BJ),

$s = 0$: reversed BJ, $s = -1$: studentized HC

GOF Tests in Terms of Local Levels

For given critical values $c_{i,n}, \tilde{c}_{i,n}, i = 1, \dots, n$, we get

$$\alpha_{i,n}^+ = \mathbb{P}(U_{i:n} \leq c_{i,n}) \text{ and } \alpha_{i,n} = \mathbb{P}(U_{i:n} \leq c_{i,n}) + \mathbb{P}(U_{i:n} \geq \tilde{c}_{i,n}).$$

$U_{i:n} \sim \text{Beta}(i, n - i + 1)$ with cdf $F_{i,n-i+1}$

One-sided test: $c_{i,n} = F_{i,n-i+1}^{-1}(\alpha_{i,n}^+), i = 1, \dots, n$

Two-sided test: Split $\alpha_{i,n}$ in $\alpha_{i,n}^{(1)}$ and $\alpha_{i,n}^{(2)}$ such that

$$\alpha_{i,n} = \alpha_{i,n}^{(1)} + \alpha_{i,n}^{(2)} \text{ (e.g., } \alpha_{i,n}^{(1)} = \alpha_{i,n}^{(2)} = \alpha_{i,n}/2)$$

$$\Rightarrow c_{i,n} = F_{i,n-i+1}^{-1}(\alpha_{i,n}^{(1)}) \text{ and } \tilde{c}_{i,n} = F_{i,n-i+1}^{-1}(1 - \alpha_{i,n}^{(2)}), i = 1, \dots, n$$

GOF Test With Equal Local Levels

$$\alpha_{1,n}^+ = \dots = \alpha_{n,n}^+ = \alpha_n^{loc} \quad \text{or} \quad \alpha_{1,n} = \dots = \alpha_{n,n} = \alpha_n^{loc}$$

$$\varphi^+(\alpha_n^{loc}): c_{i,n} = F_{i,n-i+1}^{-1}(\alpha_n^{loc})$$

$$\varphi(\alpha_n^{loc}): c_{i,n} = F_{i,n-i+1}^{-1}(\alpha_n^{loc}/2) \quad \text{and} \quad \tilde{c}_{i,n} = 1 - F_{i,n-i+1}^{-1}(\alpha_n^{loc}/2)$$

$p_{i,n} = F_{i,n-i+1}(X_{i:n})$, $i = 1, \dots, n$, one-sided p -values

$$M_n^+ = \min_{i=1, \dots, n} p_{i,n} \quad \text{and} \quad M_n = 2 \min_{i=1, \dots, n} \{p_{i,n}, 1 - p_{i,n}\}$$

$$\varphi^+(\alpha_n^{loc}) = 1 \quad \text{iff} \quad M_n^+ \leq \alpha_n^{loc} \quad \text{and} \quad \varphi(\alpha_n^{loc}) = 1 \quad \text{iff} \quad M_n \leq \alpha_n^{loc}$$

$\Rightarrow \varphi^+(\alpha_n^{loc})$ and $\varphi(\alpha_n^{loc})$ are minimum p -value (minP) GOF tests.

minP Tests: Different Names and Representations

- Berk & Jones [1978,1979] introduced M_n^+ and M_n as **minimum level attained** statistics: optimality and Bahadur efficiency;
- Buja & Rolke [2006] (unpublished): minP tests in terms of bounding functions;
- Gontscharuk, Landwehr & Finner (talks at MCP 2011 and MCP 2013): GOF tests with equal local levels, HC tests;
- Aldor et al. [2013]: tests based on the new tail-sensitive simultaneous confidence bands;
- Mary & Ferrari [2014]: non-asymptotic standardization of binomial counts, HC framework;
- Preprints: calibrated KS tests in Moskovich et al., Dirichlet-based tests in Kaplan & Goldman.

Level α minP GOF Tests

Find α_n^{loc} so that $\mathbb{P}(M_n^+ \leq \alpha_n^{loc} | H_0) = \alpha$ or $\mathbb{P}(M_n \leq \alpha_n^{loc} | H_0) = \alpha$.

Let φ be an exact level α test with local levels $\alpha_{i,n}$. Then

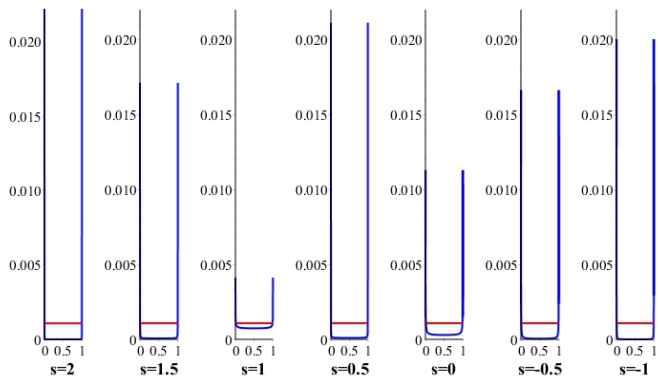
$$\min_{i=1,\dots,n} \alpha_{i,n} \leq \alpha_n^{loc} \leq \max_{i=1,\dots,n} \alpha_{i,n}.$$

Example: $i \equiv i_n$ such that $i/n \rightarrow \zeta \in (0, 1)$ leads to the asymptotic KS local level $\alpha_\zeta^{KS} = 1 - \Phi(-\log(\alpha)/(2\zeta(1 - \zeta)))$.

Hence,

$$\alpha_n^{loc} \leq 1 - \Phi(\sqrt{-2\log(\alpha)}) + o(1), \quad n \in \mathbb{N}.$$

GOF Tests Based on φ -Divergences



$\alpha_{i,n} \equiv \alpha_{i,n}(s)$ with $\alpha = 0.05$ and $n = 1000$,
 $s = 2$: Higher Criticism, $s = 1$: Berk-Jones

Higher Criticism (HC) Statistics

- HC statistics are **normalized KS statistics**:

$$HC_n^+ = \max_{i=1, \dots, n} \sqrt{n} \frac{i/n - X_{i:n}}{\sqrt{X_{i:n}(1 - X_{i:n})}},$$
$$HC_n = \max_{i=1, \dots, n} \left\{ \sqrt{n} \frac{i/n - X_{i:n}}{\sqrt{X_{i:n}(1 - X_{i:n})}}, \sqrt{n} \frac{X_{i:n} - (i-1)/n}{\sqrt{X_{i:n}(1 - X_{i:n})}} \right\}.$$

- Eicker [1979] and Jaeschke [1979] provided a lot of asymptotic results.
- Local levels of one-sided HC asymptotic level α tests

$$\alpha_{i,n}^{HC}(\alpha) \approx \frac{-\log(1 - \alpha)}{2 \log_2(n) \log(n)}$$

for the most $i \in \{1, \dots, n\}$, cf. Gontscharuk et al.[2015]

HC local levels are asymptotically equal in the sensitivity range.

Asymptotics of minP GOF Tests

Theorem 1: (Gontscharuk & Finner [2015]) The minP test with critical value d_n is an asymptotic level α test iff

$$\lim_{n \rightarrow \infty} d_n / \alpha_n^* = 1 \quad \text{with} \quad \alpha_n^* \equiv \alpha_n^*(\alpha) = -\frac{\log(1 - \alpha)}{2 \log_2(n) \log(n)}.$$

Remark:

- Critical values d_n related to asymptotic level α minP tests converge to 0 for $n \rightarrow \infty$.
- The asymptotic critical value is the same for one- and two-sided minP tests.

z -Transformed minP Statistics

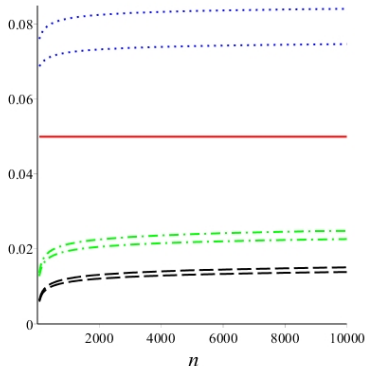
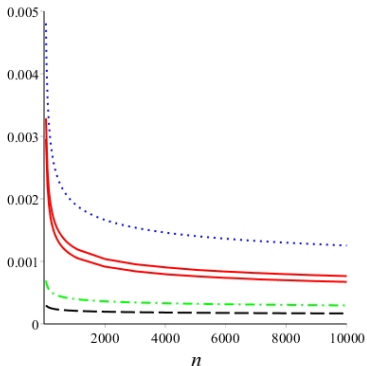
Theorem 2: (Gontscharuk & Finner [2015]) The z -transformed minP statistics $\Phi^{-1}(1 - M_n^+)$ and $\Phi^{-1}(1 - M_n/2)$ have the same asymptotic distribution as HC_n^+ and HC_n , resp.

Remark: Theorem 2 implies that the minP critical values

$$\alpha'_n = 1 - \Phi(b_n(t_\alpha^+)) \quad \text{and} \quad \alpha''_n = 2(1 - \Phi(b_n(t_\alpha))),$$

where $b_n(t) = \sqrt{2 \log_2(n)} + (\log_3(n) - \log(\pi) + 2t)/(2\sqrt{2 \log_2(n)})$, $t_\alpha^+ = -\log(-\log(1 - \alpha))$ and $t_\alpha = -\log(-\log(1 - \alpha)/2)$, lead to asymptotic level α minP tests.

Applicability of Asymptotic Results



Left graph: α_n^* , α_n' , α_n'' , α_n^{loc} related to level α one-sided (upper curves) and two-sided (lower curves) minP tests.

Right graph: $\mathbb{P}(M_n^+ \leq d_n | H_0)$ (lower curves) and $\mathbb{P}(M_n \leq d_n | H_0)$ (upper curves) for $d_n = \alpha_n^*$, α_n' , α_n''

Finite Approximation for α_n^{loc}

The asymptotic HC local levels are given by

$$\alpha_{i,n}^{HC}(\alpha) = \frac{-\log(1-\alpha)}{2\log_2(n)\log(n)} \left[1 + O\left(\frac{\log_3(n)}{\log_2(n)}\right) \right]$$

for $\log(n) \leq i \leq n - \log(n)$, cf. Gontscharuk et al. [2015].

Define

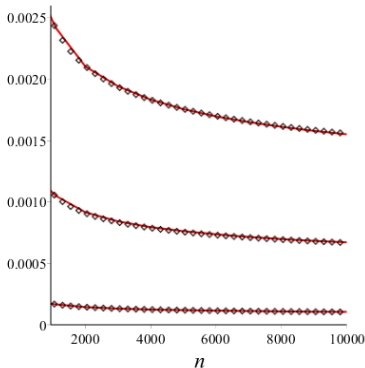
$$d_n(\alpha) = \frac{-\log(1-\alpha)}{2\log_2(n)\log(n)} \left[1 - c_\alpha \frac{\log_3(n)}{\log_2(n)} \right],$$

where $c_\alpha \in \mathbb{R}$ is a suitable constant.

Since $d_n(\alpha)/\alpha_n^* \rightarrow 1$ as $n \rightarrow \infty$, Theorem 1 implies that $d_n(\alpha)$ leads to asymptotic level α minP tests.

Approximated And Exact Critical Values

Critical values related to two-sided level α minP GOF tests



$d_n(\alpha)$ with $c_\alpha = 1.6, 1.3, 1.1$ (diamonds from bottom to top) and α_n^{loc} for $\alpha = 0.01, 0.05, 0.1$ (solid curves from bottom to top)

Simulated Global Levels Related to $d_n(\alpha)$

| n | $\alpha = 0.01$ | $\alpha = 0.05$ | $\alpha = 0.1$ |
|---------------------|-------------------|-------------------|-------------------|
| $n = 10^4$ | 0.00966 (0.00972) | 0.04874 (0.04905) | 0.09969 (0.09937) |
| $n = 5 \times 10^4$ | 0.00961 | 0.04971 | 0.10016 |
| $n = 10^5$ | 0.01018 | 0.05019 | 0.10188 |
| $n = 5 \times 10^5$ | 0.01001 | 0.05018 | 0.10115 |
| $n = 10^6$ | 0.00973 | 0.04942 | 0.10135 |

$\mathbb{P}(M_n \leq d_n(\alpha) | H_0)$ (and $\mathbb{P}(M_n \leq \alpha_n^{loc} | H_0)$ for $n = 10^4$ only)
simulated by 10^5 repetitions, where $d_n(\alpha)$ is based on
 $c_\alpha = 1.6, 1.3, 1.1$ for $\alpha = 0.01, 0.05, 0.1$, resp.

The minP GOF test works very well at least for considered
 α - and n -values.

Take Home Message

- Local levels can be seen as a measure of local sensitivity.
- One may construct a tailored GOF test by means of local levels.
- The minP test is a test with equal local levels.
- HC asymptotics is the key to the minP asymptotics.
- We provide three competing critical values leading to the asymptotic level α minP tests.
- The minP (as well as HC) asymptotics is very slow.
- New approximation for the minP critical value that works well for finite samples and leads to asymptotic level α tests.

References

- **Berk, R. and Jones, D. (1978).** Relatively optimal combinations of test statistics. *Scand. J. Stat.*, **5**, 158–162.
- **Berk, R. and Jones, D. (1979).** Goodness-of-fit test statistics that dominate the Kolmogorov Statistics. *Z. Wahrscheinlichkeit.*, **47**, 47–59.
- **Gontscharuk, V., Landwehr, S. and Finner, H. (2015).** Goodness of fit tests in terms of local levels with special emphasis on higher criticism tests. *Bernoulli*, accepted for publication.
- **Gontscharuk, V., Landwehr, S. and Finner, H. (2015).** The intermediates take it all: asymptotics of higher criticism statistics and a powerful alternative based on equal local levels. *Biometrical J.*, **57**, 159–180.
- **Gontscharuk, V. and Finner, H. (2015).** Asymptotics of goodness-of-fit tests based on minimum p -value statistics. *Commun. Stat. A-Theor.*, accepted for publication.
- **Jager, L. and Wellner, J. (2007).** Goodness-of-fit tests via phi-divergences. *Ann. Stat.*, **35**, 2018–2053.